

An introduction to econometrics of panel data

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Seminars given at UKMA departement of finance

Seminars organisation

- ▶ Two things to know about econometrics (among others)
- ▶ Panel data
- ▶ The random effect hypothesis
- ▶ A menagerie of estimators
- ▶ How to explain the discrepancies between estimators?

First thing to know: Frisch-Waugh theorem

$$\underline{y} = \hat{a}\underline{x} + \hat{b}\underline{z} + \hat{u} \quad (\text{a})$$

where \underline{x} is a variable of interest and \underline{z} a control variable

$$p_z(\underline{y}) = p_z(\hat{a}\underline{x} + \hat{b}\underline{z} + \hat{u}) = \hat{a}p_z(\underline{x}) + \hat{b}p_z(\underline{z}) + p_z(\hat{u}) \quad (\text{b})$$

where

- ▶ $p_z(\cdot)$ is the orthogonal projection onto $\mathcal{V}(\underline{z})$
- ▶ $\mathcal{V}(\underline{z})$ is the vectorial space generated by \underline{z}

but

- ▶ $p_z(\underline{z}) = \underline{z}$
- ▶ $p_z(\hat{u}) = \underline{0}$

$$\left[\underline{y} - p_z(\underline{y}) \right] = \hat{a} \left[\underline{x} - p_z(\underline{x}) \right] + \hat{u} \quad (\text{a}) - (\text{b})$$

Example

$$\begin{aligned} \underline{y} &= \hat{a}\underline{x} + \hat{b}\underline{1} + \hat{u} \\ \left[\underline{y} - \bar{y}\underline{1} \right] &= \hat{a} \left[\underline{x} - \bar{x}\underline{1} \right] + \hat{u} \end{aligned}$$

Second thing to know: what to do when $V(\underline{u}) \neq \sigma^2 I_N$?

Two ways

- ▶ Use the best estimator i.e. the FIML estimator as OLS is neither BLUE nor the FIML estimator, but
 - ▶ BLUE is not always a strong property
 - ▶ FIML estimator is the best only asymptotically
 - ▶ FIML estimator is tedious to compute
- ▶ Use OLS but do not use the “wrong” formulas
 - ▶ $\hat{a} = (X'X)^{-1}X'y$
 - ▶ $V(\hat{a}) = (X'X)^{-1}X'\Omega X(X'X)^{-1}$ if $V(\underline{u}) = \Omega$
 - ▶ Get $\widehat{\Omega}$ an convergent estimator of Ω ,
 - ▶ $\widehat{V}(\hat{a}) = (X'X)^{-1}X'\widehat{\Omega}X(X'X)^{-1}$ and the tests are valid

Panel data

- ▶ Cross section data and the limits of cross section data
- ▶ Panel data specificities
 - ▶ Individuals observed more than one time – repeated cross section
 - ▶ Time series from more than one individual – pooled time series
 - ▶ Double subscript: y_{it}, x_{it}, \dots $i = 1, \dots, N$ and $t = 1, \dots, T$
- ▶ Different kinds of panel data
 - ▶ The cross dimension dominates
 - ▶ From survey: the same households are surveyed many times
 - ▶ From administrative data: social contributions of employers, ...
 - ▶ The time dimension dominates
 - ▶ Pooled data of countries
 - ▶ Pooled high frequencies financial data

The random effect hypothesis - I: Balestra and Nerlove 1966

Basic regression model

$$y_{it} = a x_{it} + b + u_{it}$$

- ▶ Estimate an average behaviour
- ▶ Taking account individual specificities

As always, random is the preferred way to capture ignorance

- ▶ On time series, the usual way to capture persistence is the “AR(1) hypothesis”

$$u_t = \rho u_{t-1} + \varepsilon_t$$

- ▶ On panel data, the usual way to capture individual specificities is the “random effect hypothesis”

$$u_{it} = \mu_i + v_{it}$$

Data layout

Matrix notations for the basic linear model (vectors are underlined)

$$\underline{y} = X\underline{a} + \underline{u}$$

Observations are sorted by individual then by period

$$\underline{z} = \begin{pmatrix} \underline{z}_1 \\ \underline{z}_2 \\ \vdots \\ \underline{z}_i \\ \vdots \\ \underline{z}_N \end{pmatrix} \quad \underline{z}_i = \begin{pmatrix} z_{i1} \\ z_{i2} \\ \vdots \\ z_{it} \\ \vdots \\ z_{iT} \end{pmatrix}$$

The random effect hypothesis - II

Let us suppose

- ▶ no dependencies between individuals
- ▶ usual homoscedasticity hypothesis

Then

$$E(\mu_i) = 0 \quad \forall i$$

$$E(v_{it}) = 0 \quad \forall i, t$$

$$V(\mu_i) = \sigma_\mu^2 \quad \forall i$$

$$V(v_{it}) = \sigma_v^2 \quad \forall i, t$$

$$\text{Cov}(\mu_i, \mu_j) = 0 \quad \forall i \neq j$$

$$\text{Cov}(v_{it}, v_{j\tau}) = 0 \quad \forall i \neq j, t \neq \tau$$

$$\text{Cov}(\mu_i, v_{jt}) = 0 \quad \forall i, j, t$$

$$V(\underline{u}_i) = \Sigma \quad \forall i \quad \text{and} \quad E(\underline{u}_i \underline{u}_j') = 0_T \quad \forall i \neq j$$

Shape of variance-covariance matrix of \underline{u}

The variance-covariance matrix is bloc-diagonal

$$V(\underline{u}) = \begin{pmatrix} \Sigma & 0_T & \cdots & 0_T \\ 0_T & \Sigma & \cdots & 0_T \\ \vdots & \vdots & \ddots & \vdots \\ 0_T & 0_T & \cdots & \Sigma \end{pmatrix} = I_N \otimes \Sigma$$

with

$$\Sigma = \begin{pmatrix} \sigma_\mu^2 + \sigma_v^2 & \sigma_\mu^2 & \cdots & \sigma_\mu^2 \\ \sigma_\mu^2 & \sigma_\mu^2 + \sigma_v^2 & \cdots & \sigma_\mu^2 \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_\mu^2 & \sigma_\mu^2 & \cdots & \sigma_\mu^2 + \sigma_v^2 \end{pmatrix} = \sigma_v^2 W_T + (T\sigma_\mu^2 + \sigma_v^2)B_T$$

W et B matrix

$$\begin{cases} W_T = I_T - \frac{1}{T}J_T \\ B_T = \frac{1}{T}J_T \end{cases} \quad \text{with } J_T = \underline{1}_T \underline{1}'_T = \begin{pmatrix} 1 & 1 & \cdots & 1 \\ 1 & 1 & \cdots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & \cdots & 1 \end{pmatrix}$$

$$V(\underline{u}) = \sigma_v^2 W + (T\sigma_\mu^2 + \sigma_v^2)B$$

$$\begin{cases} W = I_N \otimes W_T \\ B = I_N \otimes B_T \end{cases}$$

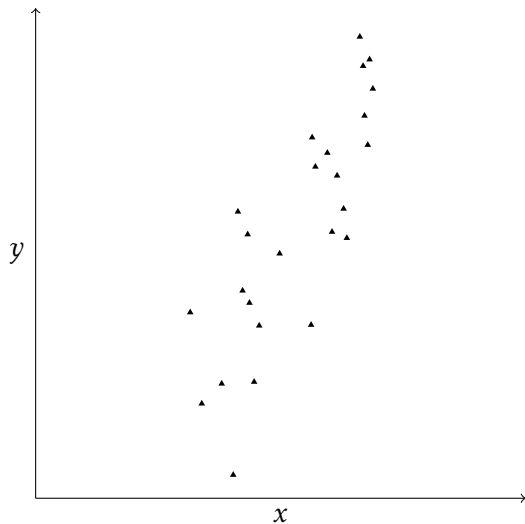
$z_{it} - \bar{z}_i$ is the element of $W\underline{z}$

\bar{z}_i is the element of $B\underline{z}$

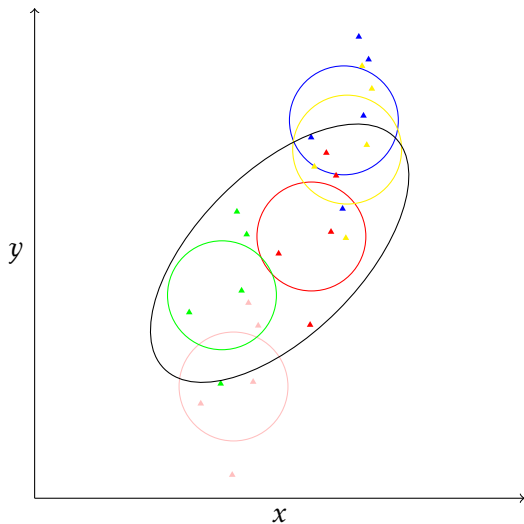
$$W^2 = W, \quad W' = W, \quad B^2 = B, \quad B' = B,$$

$$W + B = I_{NT} \quad \text{and} \quad W'B = 0_{NT}$$

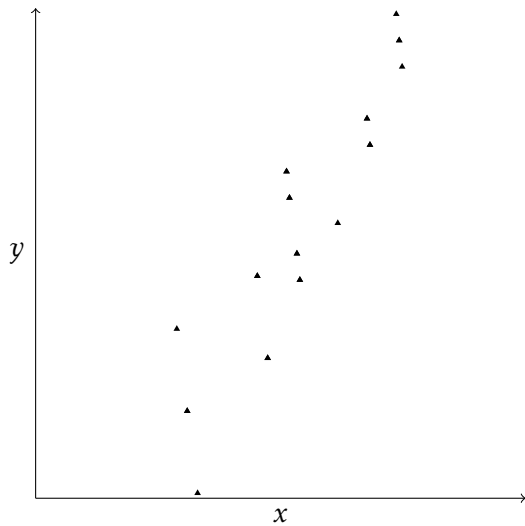
A strong link between x and y ...



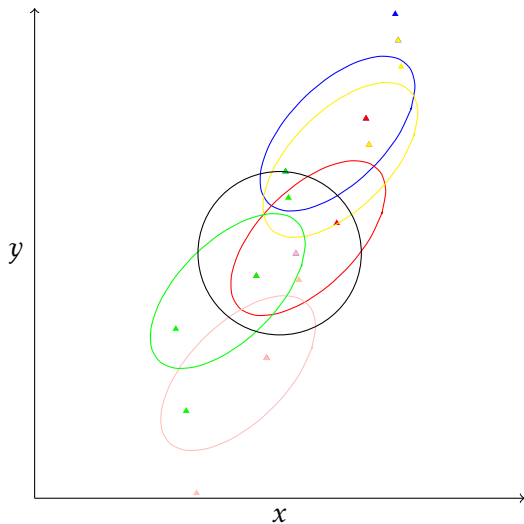
...obtained from *between* correlation



Another strong link between x and y



...obtained from *within* correlation



Application - I

Estimation on a production function on sectoral data from French national accounts

Departure equation, Cobb-Douglas specification

$$Q = AK^\alpha L^\beta e^{gt}$$

Estimated equation

$$\ln(q_{it}) = \alpha \ln(k_{it}) + \beta \ln(\ell_{it}) + gt + a + u_{it}$$